## Homework 4

Type up answers in $\mathrm{ET}_{\mathrm{E}} \mathrm{X}$ and complete according to your personal schedule. Write pseudocode in $\mathrm{ET}_{\mathrm{E}} \mathrm{X}$ and embed (e.g., R, Python, Julia) code within the same document.

1. Normal-normal Consider the setup

$$
\begin{aligned}
& \mathbf{y}=y_{1}, \ldots, y_{100} \stackrel{i i d}{\sim} \\
& \mu \operatorname{Normal}\left(\mu, \sigma^{2}\right) \\
& \sigma^{2} \sim \operatorname{Normal}\left(\mu_{0}=0, \sigma_{0}^{2}=100\right) \\
& \text { InverseGamma }(\alpha=2, \beta=10),
\end{aligned}
$$

where the second and third lines denote prior specifications, and $\mu$ and $\sigma^{2}$ are assumed to be independent a priori.
After simulating $y_{1}, \ldots, y_{100}$ with $\mu=50$ and $\sigma^{2}=10$ fixed, sample the joint posterior distribution $p\left(\mu, \sigma^{2} \mid y_{1}, \ldots, y_{100}\right)$ using
(a) Metropolis-Hastings with a bivariate Gaussian proposal distribution truncated along the dimension corresponding to $\sigma^{2}$; and
(b) Hamiltonian Monte Carlo over the transformed parameter vector $\left(\mu, \log \sigma^{2}\right)$ using 100 leapfrog iterations at each step. Note that this will require a change of variables that must be taken into account in both the posterior distribution and its gradient.

Tune both samplers (i.e., M-H proposal variance and HMC leapfrog stepsize) to average a $50 \%$ acceptance rate. Run both samplers to obtain a minimum effective sample size (ESS) of 10,000 for either parameter. What is the average ESS per sample for each sampler? Create 2 D contour plots for each of the 2 posteriors.

## 2. Hierarchically modeling house prices

(a) Start by simulating home prices in 5 fictional states. Set the true global mean to be $\mu_{0}=13.37$ and sample $j=1, \ldots, 5$ state house true price means according to the distribution $\mu_{j} \sim \mathrm{~N}\left(\mu_{0}, 4^{2}\right)$. Let each state share the same true $\sigma_{j}=\sqrt{5.35}$. Next, sample $n_{j}=100$ house prices from each state according to the distribution $y_{1 j}, \ldots, y_{n_{j} j} \stackrel{i i d}{\sim} \mathrm{~N}\left(\mu_{j}, \sigma_{j}^{2}\right)$.
(b) Next, define the hierarchical model

$$
\begin{aligned}
& y_{i j} \stackrel{i n d}{\sim} \mathrm{~N}\left(\mu_{j}, \sigma^{2}\right) \\
& \mu_{j} \stackrel{i n d}{\sim} \mathrm{~N}\left(\mu_{0}, 25^{2}\right) \\
& \mu_{0} \sim \mathrm{~N}\left(0,50^{2}\right) \\
& \sigma^{2} \sim \operatorname{Inv}-\chi^{2}(1,0.5)
\end{aligned}
$$

and build a Gibbs sampler that uses conditional conjugacy to generate closed-form updates for
i. each $\mu_{j}$ given $\mu_{0}, \sigma^{2}$ and all $y_{i j}$;
ii. $\mu_{0}$ given all $\mu_{j}$; and
iii. $\sigma^{2}$ given all $\mu_{j}$ and all $y_{i j}$.
(c) Simulate the Gibbs sampler for 10,000 iterations. Obtain $95 \%$ credible intervals for each model parameter by taking the 2.5 th and 97.5 th percentiles. Plot the 7 credible intervals along with the true values.
3. Rat tumors An experiment features $J=71$ groups of rats, each of which receive different dosages of a certain treatment. Each group features a different number of rats $n_{j}$. Denote the number of rats in each group that develop a tumor $y_{j} \leq n_{j}$. The data are available here: https://ucla-biostats-202c.github.io/code/ratTumor.txt.
(a) Define the hierarchical model

$$
\begin{gathered}
y_{j} \stackrel{i n d}{\sim} \operatorname{Binomial}\left(n_{j}, \theta_{j}\right) \\
\theta_{j} \stackrel{i i d}{\sim} \operatorname{Beta}(2, \beta) \\
\beta \sim \operatorname{Gamma}(2,1)
\end{gathered}
$$

and build a Gibbs sampler that uses
i. conditional conjugacy to generate closed-form updates for each $\theta_{j}$ given $n_{j}, y_{j}$ and $\beta$; and
ii. Metropolis-Hastings to update $\beta$ conditioned on all $71 \theta_{j}$ s. Note that you should tune the $\beta$ update step to obtain a roughly $50 \%$ acceptance rate. We sometimes refer to this scheme as "Metropolis-within-Gibbs".
(b) Simulate your Gibbs sampler for 100,000 iterations and
i. report the mean ESS for the $\theta \mathrm{s}$ as well as the ESS for $\beta$;
ii. create a figure containing
A. a $95 \%$ credible interval for each $\theta_{j} \mathrm{~s}$;
B. the posterior mean for each $\theta_{j}$;
C. the maximum likelihood estimate of each $\theta_{j}$, i.e., $\hat{\theta}_{j}=y_{j} / n_{j}$.
(c) Based on the relationship between the posterior means and the maximum likelihood estimates, guess the meaning of the phrase "Bayesian shrinkage".
4. Snoring and heart disease The following is one example of generalized linear regression within the Bayesian paradigm. We model the association between snoring and heart disease using the data in Table 1. For each individual $i$, we model the binary outcome $y_{i} \in\{0,1\}$, where $y_{i}=0$ denotes no heart disease and $y_{i}=1$ denotes heart disease. Define our vector of covariates $\mathbf{x}_{i}=\left(1, x_{i}\right)^{T}$, with first element corresponding to the intercept and $x_{i}$ corresponding to the snoring level of the $i$ th individual. Letting $\eta_{i}=\mathbf{x}_{i}^{T} \boldsymbol{\beta}$ for $\boldsymbol{\beta}=\left(\beta_{0}, \beta_{1}\right)^{T}$ the corresponding vector of regression coefficients, we have the logistic-regression likelihood

$$
\operatorname{Pr}(\mathbf{y} \mid \boldsymbol{\beta}) \propto \prod_{i}\left(\frac{e^{\eta_{i}}}{1+e^{\eta_{i}}}\right)^{y_{i}}\left(\frac{1}{1+e^{\eta_{i}}}\right)^{1-y_{i}}
$$

and the log-likelihood

$$
\begin{aligned}
\log \operatorname{Pr}(\mathbf{y} \mid \boldsymbol{\beta}) & \propto \sum_{i} y_{i} \eta_{i}-\log \left(1+e^{\eta_{i}}\right) \\
& =\sum_{j \in\{0,2,4,5\}} y_{j} \eta_{j}-n_{j} \log \left(1+\eta_{j}\right),
\end{aligned}
$$

where $y_{j}$ are the group heart disease totals and $n_{j}$ are the group sizes. If we place independent normal priors on the coefficients

$$
\beta_{0}, \beta_{1} \stackrel{i i d}{\sim} \mathrm{~N}\left(0,10^{2}\right),
$$

then their joint log-posterior is

$$
\log p(\boldsymbol{\beta} \mid \mathbf{y}) \propto-\frac{1}{200} \boldsymbol{\beta}^{T} \boldsymbol{\beta}+\sum_{j \in\{0,2,4,5\}} y_{j} \eta_{j}-n_{j} \log \left(1+\eta_{j}\right)
$$

(a) Use the Metropolis algorithm (tuned to a $50 \%$ acceptance rate) with bivariate Gaussian proposals to infer the joint posterior distribution of $\beta_{0}$ and $\beta_{1}$. How many MCMC iterations are required to achieve a minimum ESS of 1,000 ?

| Snoring level $\left(x_{j}\right)$ | Number of people with <br> heart disease $\left(y_{j}\right)$ | Total number <br> surveyed $\left(n_{j}\right)$ |
| :---: | :---: | :---: |
| 0 | 24 | 1355 |
| 2 | 35 | 603 |
| 4 | 21 | 192 |
| 5 | 30 | 224 |

Table 1: Data from 2484 subjects (reported by spouses). Snoring level $j=5$ is the most severe, and $j=0$ means no snoring.
(b) What are the posterior means and $95 \%$ credible intervals for $\beta_{0}$ and $\beta_{1}$ ? How do we interpret them?
(c) Are the results statistically significant? Why or why not?

